SELECTION ON OBSERVABLES

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- 1. Imbens, G. W. (2004). Nonparametric Estimation of Average Treatment Effects under Exogeneity: A Review. The Review of Economics and Statistics 86: 4-29
- 2. Rosenbaum, P. R., & Rubin, D. B. (1983). The central role of the propensity score in observational studies for causal effects. Biometrika, 70(1), 41-55.

¹Today's class is based on Alex Torgovitsky's notes. I'd like to thank him for kindly sharing them.

TOPICS OF PART I

Lec I: Selection on Observables

- 1. Potential Outcomes vs Latent Variables
- 2. Causal Inference
- 3. Selection Bias
- 4. Selection on Observables & Selection on Prop. Score

Lec II: Roy Models and LATE

- 1. The role of heterogeneity
- 2. Multiple instruments, Covariates, and Abadie's κ

Lec III: Marginal Treatment Effect

- 1. Parameters as functions of MTEs
- 2. Policy Relevant Treatment Effects

Lec IV: Extrapolations

- 1. Semi-Parametrics MTEs
- 2. Weights for Target Parameters

MOTIVATION

COUNTERFACTUAL QUESTIONS

- What would happen if a job training program were expanded? [Labor]
- What would happen to prices/welfare if two firms merged? [IO]
- What would different monetary policy do to real output? [Macro]
- What effect would this medication have on heart disease? [Biostatistics]
- What will happen to global temps if emissions decrease? [Climatology]

CAUSAL INFERENCE

- Thinking about a counterfactual requires thinking about causality
- Theory alone (might) tell us the direction of causality
 ⇒ Even when it does, it will rarely tell us the magnitude
- Causal inference uses data to address counterfactuals

POTENTIAL OUTCOME NOTATION

Also known as the Neyman-Fisher-Roy-Quandt-Rubin causal model

NOTATION

- ▶ \mathcal{D} is a mutually exclusive and exhaustive set of states ("treatments") e.g. training/no training $\mathcal{D} = \{0, 1\}$, prices $\mathcal{D} = [0, \infty)$, etc.
- For each $d \in D$ there is a potential outcome Y_d (a random variable)
- Y_d is what would have happened if the state were d
- **Observed**: the actual state, a random variable $D \in \mathcal{D}$
- Observed: outcome Y, related to potential outcomes as

$$Y = \sum_{d \in \mathcal{D}} Y_d I\{D = d\} = Y_D .$$

 $Y = Y_D$ is observed, but Y_d for $d \neq D$ are **unobserved**

WHAT DO WE WANT TO MEASURE?

- ▶ We are interested in **counterfactuals**, Y_d for $d \neq D$
- These variables capture the "what if" aspect of causality
- Since they are random variables, they can be summarized in many ways
- That is, there are many possible parameters of interest

EXAMPLE (PROGRAM EVALUATION)

- Suppose $d \in \{0, 1\}$ indicates participation in a job training program
- > Y is a scalar labor market outcome such as earnings
- If D = 1 we observe Y_1 (but not Y_0) and if D = 0 we observe Y_0
- There are many possible questions one could ask:
 - What would be average earnings if everyone were trained, i.e. $E[Y_1]$?
 - ▶ What is the average effect of the program, i.e. $E[Y_1 Y_0]$?
 - ▶ What about only for those who are trained, i.e. $E[Y_1 Y_0 | D = 1]$?

LATENT VARIABLE NOTATION

Many empirical models in economics look like a special case of:

Y = g(D, U) ,

where g is a function and U are unobservable variables

A **causal** interpretation of this model is implicitly saying:

 $Y_d = g(d, U)$ for every $d \in D$

This could impose assumptions, depending on what g and U are

WARNING

- Some are dogmatic about potential outcome vs. latent variable notation
- Often follows some field-specific social norms, e.g. labor vs. IO
- Remember: It's just notation use the above to translate

OUTLINE

1. Potential Outcomes and Latent Variables

- 2. Random Assignment: The fundamental Problem of Causal Inference
- 3. Selection Bias and Selection on Observables
- 4. The Role on the Propensity Score
- 5. Final Remarks on Selection on Observables

- **Random assignment** is the assumption that $\{Y_d : d \in D\} \perp D$
- ▶ Under random assignment, the distribution of Y_d is point identified,

$$\underbrace{F_d(y)}_{\text{parameter}} = P\{Y_d \le y\} \stackrel{(*)}{=} P\{Y_d \le y | D = d\} = \underbrace{P\{Y \le y | D = d\}}_{observed},$$

where (*) follows from random assignment.

- Any parameter that is a function of $F_d : d \in D$ is also point identified
- Intuition: conditioning on treatment does not change potential outcomes
- \Rightarrow No self-selection, sorting, correlated observables/unobservables, etc.

Common parameters of interest with binary D

- Average treatment effect (ATE): $E[Y_1 Y_0]$
- Average treatment on the treated (ATT): $E[Y_1 Y_0|D = 1]$
- ▶ Quantile treatment effect (QTE): $Q_{Y_1}(t) Q_{Y_0}(t)$ for some $t \in (0, 1)$
- QTE on the treated/untreated (QTT/QTU) defined analogously
- All point identified under random assignment
- Moreover, ATE = ATT = ATU, and QTE = QTT = QTU
- Nothing systematically different about treatment/control groups
- ▶ If *D* is multivalued or continuous, conventions are less established

CAUSAL INFERENCE: PROBLEM

THE PROBLEM

- Even with random assignment, joint dist's, $P\{Y_1 \le y_1, Y_0 \le y_0\}$, are not point id'd:
- Sometimes called the fundamental problem of causal inference
- ▶ Intuition: we never see both *Y*⁰ and *Y*¹ for anyone

IMPLICATIONS

- Most features of $Y_1 Y_0$ are not point identified
- Even with random assignment \Rightarrow so without it as well
- We might care about the proportion of individuals who are hurt:

 $P{Y_1 \leq Y_0} \rightarrow \text{ not point identified!}$

- ▶ Nor are the quantiles of $Y_1 Y_0$
- Note: Quantile treatment effect (QTE) vs. quantile of the treatment effect

RANDOM ASSIGNMENT AND COVARIATES

ROLE OF COVARIATES

- Suppose we regress *Y* on *D* and predetermined *X*
- ▶ D is randomly assigned, so should be uncorrelated with $X \rightarrow$ Common practice to check this as a "balance test"
- Also means variation in coefficient on D will go down
- How much depends on how much X and Y are correlated

EXAMPLE

- > Y is cholesterol after the experiment
- D is a drug intended to reduce cholesterol
- > X is your cholesterol in the past, before the experiment
- X probably explains a lot of the variation in Y
- Controlling for X reduces residual variation in Y, but not D
- This allows one to estimate the effect of D more precisely

SCOPE OF RANDOM ASSIGNMENT

WHEN IS RANDOM ASSIGNMENT A GOOD ASSUMPTION?

- Typically, settings where agents have no control over D
- Less likely: Agents choose D without considering $\{Y_d : d \in D\}$
- Randomized controlled experiments are the leading case
- Random assignment is rarely compelling with observational data
- When agents can control D, we typically expect selection
- Random assignment leads to high internal validity
- The phrase "gold standard" is often used in biostatistics
- In practice, researchers rarely "flip a coin" (c.f. CAR)
- Random assignment often comes along with lower external validity

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SELECTION

FORMAL DEFINITION

▶ There is selection into the treatment state *D* if

 $Y_d|D = d$ is distributed **differently** from $Y_d|D = d'$ for $d \neq d'$

Expected to occur if agents choose D with knowledge of $\{Y_d : d \in D\}$

SELECTION IS COMMON

- Particularly concerning if you are trained in neoclassical economics
- Optimization: Agents choose a job training program $(D \in \{0,1\})$ to maximize utility
- Utility: incorporates expected future earnings (Y_0, Y_1)
- Agents who choose job training might do so because of low Υ₀
- Alternatively, might choose D = 0 because of high Y_0

SELECTION BIAS

- Consider the simple treatment/control mean contrast under selection
- This contrast would be the ATE under random assignment
- Decompose the contrast into a causal effect and selection bias:

$$E[Y|D = 1] - E[Y|D = 0] = \underbrace{(E[Y_1|D = 1] - E[Y_0|D = 1])}_{ATT} + \underbrace{(E[Y_0|D = 1] - E[Y_0|D = 0])}_{\text{selection bias}}$$

- First term: causal effect for those who were treated
- Second term: how the treated would have been different anyway
- ► Effects could cancel out: ATT is (+) while selection bias is (-) ⇒ Job training program, drug for a lethal disease, etc

- A simple relaxation of random assignment is selection on observables
- Suppose that we observe (Y, D, X) where X are covariates
- The selection on observables assumption is that

 $\{Y_d: d\in \mathcal{D}\} \perp D \mid X.$

- Says: Conditional on *X*, treatment is as-good-as randomly assigned
- Other terms: unconfoundedness, ignorable treatment assignment
- Underlies causal interpretations of linear regression

Conditional version of random assignment:

$$\begin{aligned} F_d(y|x) &= P\{Y_d \leq y|X=x\} \\ &= P\{Y_d \leq y|D=d, X=x\} \\ &= P\{Y \leq y|D=d, X=x\} \end{aligned}$$

- Second equality requires overlap: $P\{D = d | X = x\} > 0$
- Integrating over x, one can point identify the marginals

$$F_d(y) = P\{Y_d \le y\} = E[P\{Y_d \le y | X = x\}] = E[P\{Y \le y | D = d, X = x\}]$$

IDENTIFICATION OF MEAN CONTRASTS

- Suppose $D \in \{0,1\}$ is binary by far the most common case
- Using essentially the same argument as on the previous page:

 $ATE = E[E[Y_1|X] - E[Y_0|X]] = E[E[Y|D = 1, X] - E[Y|D = 0, X]]$

Similar expression for the ATT has an important difference:

$$ATT = E[Y|D = 1] - E[E[Y|D = 0, X|D = 1]]$$

- Helps in estimation since only one conditional expectation (more later)
- Note that only mean independence is needed for these arguments

$$E[Y_d|D = 0, X] = E[Y_d|D = 1, X]$$

Difficult to think of arguments for means (without full) independence

EXAMPLE

- Suppose that $(Y_0, Y_1) \perp D | X$
- Let X₂ be a subset of X
- Let X₁ be a subset of X₂
- So controlling on X is the most information, and X₁ is the least
- Suppose however that we only have X₂ (hence X₁) in our data
- Perhaps surprisingly, using X₂ can introduce more bias than using X₁
- ▶ That is: adding information (X₂ vs X₁) **need not** reduce bias
- ▶ If *X*₂ = *X*, then it does, but not more generally
- Point is not well-appreciated in applied work. But should be concerning

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THE PROPENSITY SCORE

DEFINITION

- **b** Binary treatment case: $D \in \{0, 1\}$
- ▶ $p(x) = P{D = 1 | X = x}$ is called the **propensity score**
- Let P = p(X) be the random variable $P\{D = 1|X\}$

ROSENBAUM AND RUBIN (1983) SUFFICIENCY ARGUMENT

▶ Selection on observables implies $(Y_0, Y_1) \perp D \mid P$:

$$P\{D = 1 | Y_0, Y_1, P\} = E \Big[P\{D = 1 | Y_0, Y_1, P, X\} | Y_0, Y_1, P \Big]$$
$$= E \Big[P\{D = 1 | Y_0, Y_1, X\} | Y_0, Y_1, P \Big]$$
$$= E \Big[P\{D = 1 | X\} | Y_0, Y_1, P \Big]$$
$$= E \Big[p(X) | Y_0, Y_1, P \Big] = P$$

Implication: we can condition on P instead of X

PROPENSITY SCORE WEIGHTING

Using p, we can write the ATE as a weighted average of Y

$$ATE(x) = E\left[\frac{Y(D - p(x))}{p(x)(1 - p(x))}\Big| X = x\right]$$

Average over X to point identify

$$ATE = E\left[\frac{Y(D - p(X))}{p(X)(1 - p(X))}\right]$$

Similar expressions can be derived for ATT,

$$ATT = E\left[\frac{Y(D - p(X))}{P\{D = 1\}(1 - p(X))}\right]$$

Derivations are straightforward (see Pset)

DIFFERENT IDENTIFICATION ARGUMENTS

SUMMARY

- ► Three different constructive identification results for the ATE ⇒ Match on X, match on P, weight using p
- Each one shows that ATE is point identified
- And they are all derived under the same assumptions

SO WHY HAVE THREE?

- Identification arguments directly inform the construction of estimators
- Different arguments suggest different estimators
- In general, these different estimators may have different properties
- In selection on observables this is definitely true
 Subtle differences in efficiency, rates of convergence
- More importantly, differences in finite sample performance

MULTIVALUED TREATMENTS

- Many interesting counterfactual states are multivalued
- The main identification arguments mostly remain the same/similar
- Some details (Imbens, 2000) regarding the (generalized) propensity score
- However, the literature is overwhelmingly about $D \in \{0, 1\}$
- Imbens and Rubin (2015 book, 650 pages) exclusively discuss this case!
- The reason seems (to me) to be mostly sociological
- Nonparametric methods are highly valued by those in this literature
- With $D \in 0, 1$ there is only non-parametric (in D)
- ▶ If D ∈ {0,1,2}, then one needs to make a choice:
 (a) Make a (potentially wrong) functional form assumption
 (b) Remain non-parametric basically reduces back to the binary case
- Community is against the first option, and second has low payoff for theory work.

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CRITICISMS OF SELECTION ON OBSERVABLES

INHERENT UNOBSERVABLES

- Selection on observables can be difficult to believe in economics
- ightarrow Inherent unobservables: preferences, private info, expectations, \ldots
- Observationally identical people behave differently due to . . . a coin flip?

CONTROLLING FOR MORE: NOT A SOLUTION

- Often argued that large X makes selection on observables more likely ⇒ This is, of course, not necessarily true - we saw this earlier
- ► Even if it were, still raises an uncomfortable friction with overlap ⇒ If we could perfectly explain D with X then $P\{D = 1|X\}$ would be 0 or 1

BETTER METHODS FOR CHOOSING OBSERVABLES WILL NOT SOLVE THIS

- Selection on observables is seeing a resurgence with machine learning
- Fancier methods, but the identifying assumption is still the same
- Bias/variance trade-off is not the first-order issue here

ANGRIST (1998)

MISFIT SELECTION ON OBSERVABLES

- Well-known economic application of selection on observables:
 - Y is a labor market outcome (employment/earnings)
 - ▶ $D \in \{0,1\}$ is veteran status (participation in the military)
 - X are socioeconomic variables (race, year, schooling, AFQT, age)
- Assumption: given X, military participation as-if randomly assigned
- Observationally similar people randomly join the military?!?
- Ignores first-order issues such as outside employment options
 Also unobservable screening factors (fitness, interpersonal skills)
- These are unobserved and inherently unobservable

ALLOWING FOR SELECTION ON UNOBSERVABLES

- Most applied microeconomists seem to share this skepticism
- This motivates the other methods that we will discuss in the course
- All use different arguments to allow for selection on unobservables